

Network-based Integration of Microbiome and Metabolomic Data

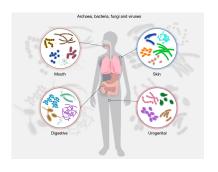
Banff, 17 July 2025

Jing Ma

The human microbiome



All of the microbes and their genome, mostly bacteria



- More microbial cells than our somatic cells
- More microbial genes than our human genome
- Compositions vary within a person and between individuals
- Highly dynamic yet robust
- Association with many diseases

Microbiome data



OTU	Species	Sample 1	Sample 2	Sample 3
1	E.coli	17	0	335
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Table 1: A typical microbiome contingency table

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high-dimensional

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- ullet compositional: only relative abundances are meaningful o normalization

Metabolomics



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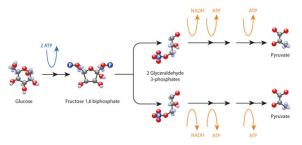


Figure 1: Glycolysis: energy is used to convert glucose to a 6 carbon form. Thereafter, energy is generated to create two molecules of pyruvate. Credit to Nature Education.

Metabolomic data



Compound	Sample 1	Sample 2	Sample 3
Glucose	42,062,493	46,507,270	48,849,105
Glutamic acid	1,027,679	1,317,161	2,527,070
Propionate	3,487	6,262	9,188

Table 2: Peak intensities of compounds across samples

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- high variance \rightarrow log transformation



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- ATP: metabolites that are synthesized de novo by gut microbes

Detecting microbial metabolites



Knowledge gap:

- most of the bacterial metabolites remain unidentified.
- many known metabolites have yet to be functionally characterized.

¹Morton et al. 19'. Nat Methods; Reiman et al. 21'. PLOS Comp Bio

²Quinn-Bohmann et al. 25'. Nat Microbiol

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Current methods for learning microbial-metabolite interactions:

- correlation networks
- machine learning models¹
- mechanistic models²

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Microbe-metabolite networks



Let $A \in \{0,1\}^{p_1 \times p_2}$ denote the latent network between p_1 microbes and p_2 metabolites.

In this bipartite network, the nodes are microbes/metabolites and the edges represent associations between microbes and metabolites.

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In this bipartite network, the nodes are microbes/metabolites and the edges represent associations between microbes and metabolites.

Inference for A amounts to testing:

$$H_{0,i,j}: A_{i,j} = 0$$
 versus. $H_{1,i,j}: A_{i,j} \neq 0$,

for all $1 \le i \le p_1, 1 \le j \le p_2$.

False discovery rate



FDR provides a way of quantifying the statistical significance of multiple hypothesis tests.

	Not significant	Significant	Total
Null is true	N_{00}	N_{10}	m_0
Alternative is true	N ₀₁	N_{11}	m_1
Total	S	R	m

Table 3: Classification of tested hypothesis

$$\label{eq:fdr} \text{FDR} = \textit{E}(\frac{\textit{N}_{10}}{\textit{R} \vee 1}), \quad \text{mFDR} = \frac{\textit{E}(\textit{N}_{10})}{\textit{E}(\textit{R})}$$

FDR control



The Benjamini & Hochberg (BH) procedure³

- Choose a desired significance level $\alpha \in (0,1)$.
- Sort the p-values in increasing order: $p_{(1)} \leq p_{(2)} \leq \cdots \leq p_{(m)}$.
- Find the largest index k such that:

$$p_{(k)} \leq \frac{k}{m} \alpha$$

• Reject all hypotheses with p-values $p_i \leq p_{(k)}$.

³Benjamini & Hochberg. (95') JRSSB

Local false discovery rate



Let $X = (x_i) \in \mathbb{R}^{p_1 p_2}$ denote the observations (e.g., z-scores). Efron et al.⁴ studied the mixture model

$$f(x) = \pi_0 f_0(x) + \pi_1 f_1(x),$$

for multiple testing, where

- f_0 : null distribution (e.g., $\mathcal{N}(0,1)$)
- f₁: the alternative distribution
- $\pi_0, \pi_1 = 1 \pi_0$: prior probabilities

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The Empirical Bayes local false discovery rate⁵ is:

$$\mathsf{IFDR}(x) = \frac{\pi_0 f_0(x)}{f(x)}.$$

Reject hypotheses with IFDR(x) < α .

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⁵Sun and Cai. (07') JASA

Incorporating structures



More power can be achieved by exploiting the local dependence structure, e.g., Hidden Markov models⁶.

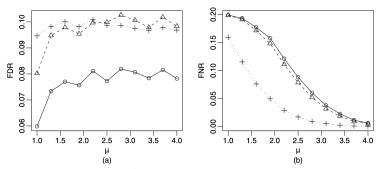


Fig. 1. Comparison of BH (\bigcirc), AP (\triangle) and OR (+) in an HMM (the FDR level is set at 0.10): (a) FDR *versus* μ ; (b) FNR *versus* μ

⁶Sun and Cai (09') JRSSB

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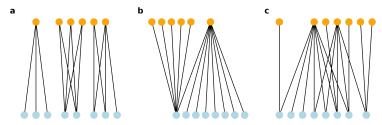


Figure 2: Topology of interest: (a) three biclusters, (b) fully nested graph, and (c) preferential attachment.

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Latent graph model



Let $X=(x_{i,j})\in\mathbb{R}^{p_1\times p_2}$ denote the matrix-valued observations (e.g., z-scores).

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Let $X = (x_{i,j}) \in \mathbb{R}^{p_1 \times p_2}$ denote the matrix-valued observations (e.g., z-scores).

We model X using a latent bipartite stochastic block model (biSBM), defined with respect to row clustering $Z_1 = (Z_{i,1})$ and column clustering $Z_2 = (Z_{j,2})$. For $i = 1, \ldots, p_1$ and $j = 1, \ldots, p_2$:

$$Z_{i,1} \sim \mathsf{Multi}(1, \alpha_1),$$
 $Z_{j,2} \sim \mathsf{Multi}(1, \alpha_2),$ $A_{i,j} \mid Z_1, Z_2 \sim \mathsf{Bern}(\pi_{Z_{i,1}, Z_{j,2}}),$ $X_{i,j} \mid A_{i,j}, Z_1, Z_2 \sim A_{i,j} g_{\nu_{Z_{i,1}, Z_{j,2}}} + (1 - A_{i,j}) g_{0,\nu_0}.$ (1)

Model parameters are $\theta = (\alpha_1, \alpha_2, \pi, \nu, \nu_0)$.

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Reject the hypotheses if the ℓ -value is small, where the threshold is chosen to control mFDR.

The *structured* ℓ *values* provide much more information than a single observation $x_{i,j}$ and will considerably help to make the final decision.

Identifiability



The *Gaussian* noisy biSBM is identifiable under the constraint that all elements of $\{(0, \sigma_0), (\mu_{q,l}, \sigma_{q,l}), 1 \le q \le B_1, 1 \le l \le B_2\}$ are distinct.

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In general, the requirement of all elements being distinct is not necessary, as the model is also identifiable if there is a single alternative distribution such that $\sigma_{q,l} = \sigma$ and $\mu_{q,l} = \mu$.

Parameter estimation



Need to use the EM algorithm due to the latent variables A, Z_1, Z_2 . Let Q denote a probability distribution of the latent variables.

$$\log \mathcal{L}(X; \boldsymbol{\theta}) = \underbrace{E_{Q}[\log \mathcal{L}(X, A, Z_{1}, Z_{2}; \boldsymbol{\theta})] + \mathcal{H}(Q)}_{ELBO} + KL(Q \| P_{A, Z_{1}, Z_{2} | X; \boldsymbol{\theta}}),$$
(2)

- **1** Initialize $\theta^{(0)}$.
- **2** E-step: evaluate the expectation in (2) with respect to $Q = P_{A,Z_1,Z_2|X;\theta^{(t)}}$.
 - Mean-field approximation
- **3** M-step: update $\theta^{(t+1)}$ by maximizing the ELBO.
- 4 Iterate between E- and M-step until convergence.

Selecting the number of clusters



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We use the integrated classification likelihood (ICL) criterion to select the optimal B_1 and B_2 , allowing them to be different.

Simulations



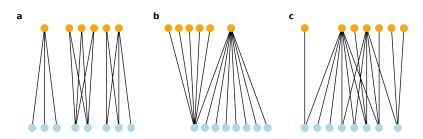


Figure 3: Illustrations of the latent bipartite network used in simulations: (a) three biclusters, (b) fully nested graph, and (c) preferential attachment

- (a) Modular structures
- (b) Nested graph in ecology: generalist vs specialist
- (c) Preferential attachment: the rich gets richer

Simulation parameters



- $p_1 = 150, p_2 = 200$
- $\mathcal{N}(0,1)$ versus $\mathcal{N}(2,1)$
- Compare the average performance over 100 simulations
- Both the new and the SC procedures⁸ were implemented assuming known null density.

⁸Sun and Cai (07') JASA



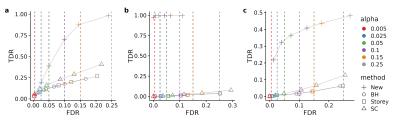


Figure 4: Plot of the empirical (FDR, TDR) as a function of the nominal level α for the new procedure, BH, Storey's q-value, and the SC procedure. Dashed lines indicate the nominal level α .



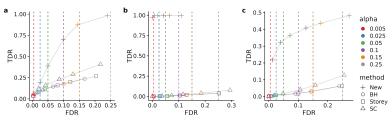


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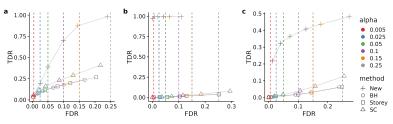


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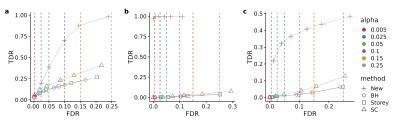


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- (b) ICL selected two biclusters in 83% of the simulations.
- (c) ICL mostly found three clusters in type I vertex and one cluster in type II vertex.

Bacterial vaginosis



- BV is the most common vaginal condition, affecting an estimated 30% of women at any given time⁹.
- BV is associated with increased transmission of HIV and STIs as well as increased risk of preterm labour.
- Diagnosis relies on microscopy to identify BV-like bacteria by morphology alone (Nugent Scoring).
- The pathogenesis of BV remains unclear.

⁹McMillan et al. (15') Scientific Reports



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- Aim is to understand the microbial functional changes during BV.



Let \hat{Y}_1 and \hat{Y}_2 denote, respectively, the standardized data. The sample correlation is defined by

$$\hat{\rho}_{i,j} = \frac{1}{n} \sum_{k=1}^{n} \hat{Y}_{1,k,i} \hat{Y}_{2,k,j}.$$

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Let

$$s_{i,j} = \frac{1}{n} \sum_{i=1}^{n} (2 \hat{Y}_{1,k,i} \hat{Y}_{2,k,j} - \hat{\rho}_{k,i} \hat{Y}_{1,k,i} - \hat{\rho}_{k,i} \hat{Y}_{2,k,j})^{2}.$$

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The test statistic

$$x_{i,j} = rac{2\hat{
ho}_{i,j}}{\sqrt{s_{i,j}/n}}
ightarrow \mathcal{N}(0,1),$$

under finite fourth moment condition¹⁰.

¹⁰Cai and Liu (16') JASA



Two-sample inference comparing BV to normal patients:

$$x_{i,j} = \frac{2(\hat{\rho}_{i,j}^{(1)} - \hat{\rho}_{i,j}^{(2)})}{\sqrt{s_{i,j}^{(1)}/n_1 + s_{i,j}^{(2)}/n_2}}$$

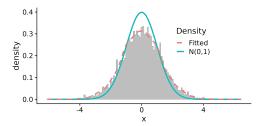


Figure 5: Histogram of observed *z*-scores compared to the standard normal and the estimated marginal distribution by the proposed approach.

Results



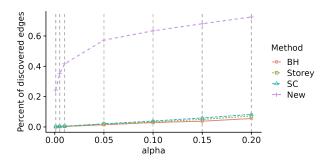


Figure 6: Percent of rejected edges as a function of the significance level α for the different procedures.

Results



The new procedure groups microbes and metabolites with similar association patterns into biclusters.

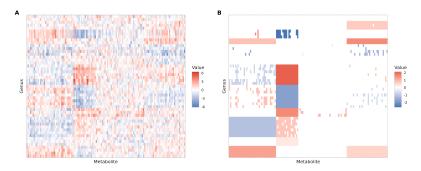


Figure 7: Heat map of the data (A) compared to the estimated graph by the proposed approach at $\alpha=0.1\%$ (B). Rows and columns are ordered by the inferred clustering.

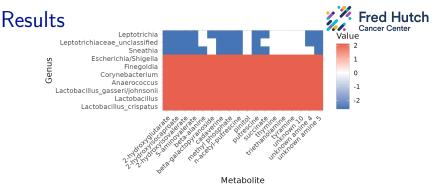


Figure 8: A zoom in view of two biclusters with the largest mean difference.

- Top bicluster consists of Leptotrichia and Sneathia which are emerging pathogens implicated in BV. Association of these genera with metabolites is higher in BV patients.
- Bottom bicluster consists of Lactobacillus species, important for keeping a healthy vaginal microbiome. Association of these genera with metabolites is higher in normal individuals.
- Shed light on uncharacterized metabolites through "Guilt by Association".

Summary



metaMint

preprint: arXiv.2506.12275

Software: https://github.com/drjingma/metaMint

Future directions

Degree heterogeneity is not accounted for.

Computation: the method is slower than existing methods.
 Alternative model estimation and/or selection is helpful.

