

Variance Components Estimation for Linear Mixed Models

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Heritability Estimation



The ACE model

$$y = \mathbf{X}\beta + \gamma + \mathbf{c} + \mathbf{e}$$

$$\gamma \sim N(0, \sigma_{\gamma}^{2} \mathbf{A}), \quad \mathbf{c} \sim N(0, \sigma_{c}^{2} \mathbf{C}), \quad \mathbf{e} \sim N(0, \sigma_{e}^{2} I_{n})$$

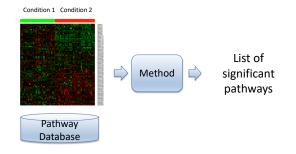
- y: $n \times 1$ vector of quantitative traits
- A: $n \times n$ genetic related matrix
- C: $n \times n$ matrix for shared environment

$$h^2 = \sigma_\gamma^2 / (\sigma_\gamma^2 + \sigma_c^2 + \sigma_e^2)$$



Scientific Question

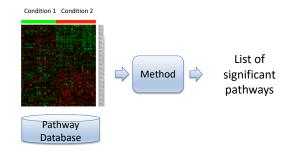
Whether a *genetic/metabolic pathway* is involved in responding to changes in environmental conditions or in specific cell functions.





Scientific Question

Whether a *genetic/metabolic pathway* is involved in responding to changes in environmental conditions or in specific cell functions.



Reduce the complexity; more explanatory power.

Pathway Enrichment Analysis



The NetGSA model¹ (for $t = 1, 2, i = 1, \ldots, n_t$)

$$y_i^{(t)} = \frac{\mu^{(t)} + \gamma_i^{(t)}}{\gamma^{(t)}} + e_i^{(t)}$$
$$\gamma^{(t)} \sim \mathcal{N}(0, A_t)$$
$$e_i^{(t)} \sim \mathcal{N}(0, \sigma_e^2 I_p)$$

• $y_i^{(t)}$: $p \times 1$ vector of observation for individual *i* in group *t*

- A_t^{-1} : $p \times p$ network information matrix
- Test statistic for $H_0: \mu_G^{(1)} = \mu_G^{(2)}$ depends on the variance components.

¹ Ma J, et al. Bioinformatics. 2016

VC Estimation with REML



Pros

Statistically efficient

Cons

- ► Need to invert n × n matrices (or p × p matrices in NetGSA) → computationally expensive, e.g.
 - n > 100K in heritability estimation
 - $p \approx 3K$ in enrichment analysis

Haseman-Elston Regression²



The residual after removing fixed effects is

$$\varepsilon = \mathbf{y} - \mathbf{X}\beta = \gamma + \mathbf{c} + \mathbf{e},$$

whose second moment is

$$\mathbb{E}(\varepsilon\varepsilon') = \sigma_{\gamma}^{2} \mathbf{A} + \sigma_{c}^{2} \mathbf{C} + \sigma_{e}^{2} \mathbf{I}_{n}$$

² Sofer T. Stat. Appl. Genet. Mol. Biol. 2017

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Let Vec(A) denote the upper triangular part of a matrix A including the diagonal. The HE method solves for σ_i^2 by regressing

$$\widetilde{Y} = \operatorname{Vec}(\widehat{\varepsilon}\widehat{\varepsilon}') \in \mathbb{R}^{n^*}, \quad n^* = \frac{n(n+1)}{2}$$

on the design matrix

$$\widetilde{X} = [\operatorname{Vec}(I_n), \operatorname{Vec}(A), \operatorname{Vec}(C)] \in \mathbb{R}^{n^* \times 3}$$

² Sofer T. Stat. Appl. Genet. Mol. Biol. 2017



Pros

Only need to invert a 3 × 3 matrix

Cons

- May get negative estimates
- ► Computational cost is $O(n^2)$ in the ACE model and $O(np^2)$ in NetGSA \rightarrow can be inefficient if *n* and *p* are large



Negative estimates

↓ Use non-negative least squares (NNLS)

NNLS



NNLS solves for the variance components by minimizing

$$\frac{1}{n^*}\left\{\theta'\widetilde{X}'\widetilde{X}\theta-2\theta'\widetilde{X}'\widetilde{Y}\right\}, \quad s.t. \ \theta \geq 0,$$

where

$$\frac{1}{n^*}\widetilde{X}'\widetilde{X} = \frac{1}{n^*}\sum_{i=1}^{n^*}\widetilde{X}'_i\widetilde{X}_i,$$
$$\frac{1}{n^*}\widetilde{X}'\widetilde{Y} = \frac{1}{n^*}\sum_{i=1}^{n^*}\widetilde{X}'_i\widetilde{Y}_i.$$



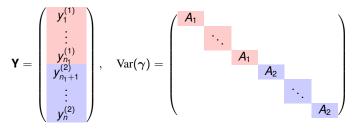
- NNLS depends only on the average products between rows of X and between rows of X and Y.
- We can approximate these values by subsampling rows of X and Y
 multiple times to get robust NNLS estimates.

Subsampling for NetGSA



 $\mathbf{Y} = \boldsymbol{\mu} + \boldsymbol{\gamma} + \boldsymbol{\varepsilon}$

where



• $Var(\gamma)$ has $n = n_1 + n_2$ diagonal block.

We can subsample observations and genes simultaneously!

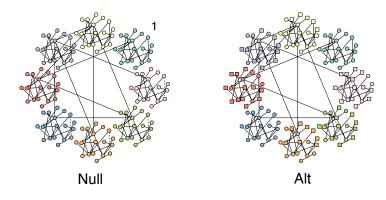


- Classical CLT fails because entries in \widetilde{Y} are dependent
- We evoke the CLT for weakly dependent processes to conclude consistency for the HE/REHE estimator

Simulations: Setup



8 subnetworks with varying degrees of enrichment

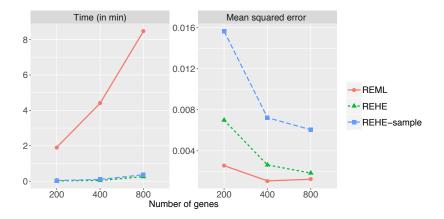


● set 1 ● set 2 ● set 3 ● set 4 ● set 5 ● set 6 ● set 7 ● set 8

Simulations: VC Estimation



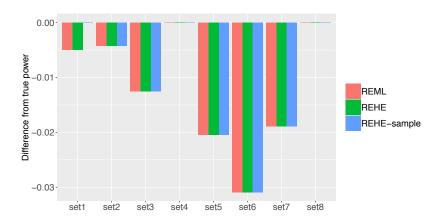
•
$$n_1 = n_2 = 200, \sigma_{\gamma}^2 = \sigma_{\epsilon}^2 = 1$$



Simulations: Power



•
$$p = 800, n_1 = n_2 = 200, \sigma_{\gamma}^2 = \sigma_{\epsilon}^2 = 1, \mu^{(2)} - \mu^{(1)} = 0.1$$



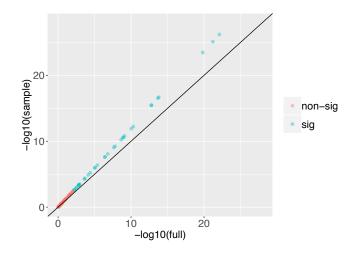


- Gene expression from 160 normal vs 264 tumor samples
- p = 2800 genes with Entrez IDs
- Network topology information extracted from BioGrid
- Analysis of 96 KEGG signaling pathways

TCGA Prostate Cancer Pathway Analysis



Each dot represents one pathway; FDR p-value threshold at 0.01



Extensions



The ACE model with $\mathbf{X} \in \mathbb{R}^{n \times m}$

$$y = \mathbf{X}\beta + \gamma + \mathbf{c} + \mathbf{e}.$$

Let *L* be the $(n - m) \times n$ matrix with its rows spanning the kernel space of **X**'. Then

$$\mathbb{E}[L \mathbf{y}\mathbf{y}' \mathbf{L}'] = L(\sigma_{\gamma}^2 \mathbf{A} + \sigma_c^2 \mathbf{C} + \sigma_e^2 \mathbf{I}_n)\mathbf{L}'.$$

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> yy' is a sample outer product derived from the Euclidean distance.

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Let *L* be the $(n - m) \times n$ matrix with its rows spanning the kernel space of **X**'. Then

$$\mathbb{E}[L \mathbf{y}\mathbf{y}' L'] = L(\sigma_{\gamma}^2 \mathbf{A} + \sigma_c^2 C + \sigma_e^2 I_n)L'.$$

- > yy' is a sample outer product derived from the Euclidean distance.
- If we do not observe y but have an outer product matrix M defined from a distance measure suitable for microbiome data, we can detect heritable microbial communities by

$$\mathbb{E}[L \ \mathbf{M} \ \mathbf{L}'] = L(\sigma_{\gamma}^{2}\mathbf{A} + \sigma_{c}^{2}\mathbf{C} + \sigma_{e}^{2}\mathbf{I}_{n})\mathbf{L}'.$$

Collaborators



- ► Kun Yue (UW)
- Ali Shojaie (UW)