

**SUPPLEMENT TO "GENERALIZED MATRIX DECOMPOSITION
REGRESSION: ESTIMATION AND INFERENCE FOR TWO-WAY
STRUCTURED DATA"**

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In this supplement, we prove the theoretical results in the main paper.

1. Derivations of Eq. (7) in the main text. We only focus on the high-dimensional case where $K = n < p$, noting that similar derivations can be straightforwardly applied to the case where $K > p$. It follows from the definition of $\widehat{\beta}_{\text{KPR}}(\eta)$ that

$$\begin{aligned}\widehat{\beta}_{\text{KPR}}(\eta) &= (\mathbf{X}^\top \mathbf{H} \mathbf{X} + \eta \mathbf{Q}^{-1})^{-1} \mathbf{X}^\top \mathbf{H} \mathbf{y} \\ &= \mathbf{Q} (\mathbf{X}^\top \mathbf{H} \mathbf{X} \mathbf{Q} + \eta \mathbf{I}_p)^{-1} \mathbf{X}^\top \mathbf{H} \mathbf{y} \\ &= \mathbf{Q} \mathbf{X}^\top \mathbf{H} (\mathbf{X} \mathbf{Q} \mathbf{X}^\top \mathbf{H} + \eta \mathbf{I}_n)^{-1} \mathbf{y}.\end{aligned}$$

Since the GMD of \mathbf{X} yields $\mathbf{X} = \mathbf{U} \mathbf{S} \mathbf{V}^\top$, we get

$$\begin{aligned}\widehat{\beta}_{\text{KPR}}(\eta) &= \mathbf{Q} \mathbf{V} \mathbf{S} \mathbf{U}^\top \mathbf{H} (\mathbf{U} (\mathbf{S}^2 + \eta \mathbf{I}_n) \mathbf{U}^\top \mathbf{H})^{-1} \mathbf{y} \\ (1) \quad &= \mathbf{Q} \mathbf{V} \mathbf{S}^{-1} \mathbf{S}^2 (\mathbf{S}^2 + \eta \mathbf{I}_n)^{-1} \mathbf{U}^\top \mathbf{H} \mathbf{y}.\end{aligned}$$

The last equality in (1) comes from the fact that \mathbf{U} is a $n \times n$ invertible matrix and $\mathbf{U}^\top \mathbf{H} \mathbf{U} = \mathbf{I}_n$. Denoting $\mathbf{W}_\eta = \mathbf{S}^2 (\mathbf{S}^2 + \eta \mathbf{I}_n)^{-1}$, we can write

$$(2) \quad \widehat{\beta}_{\text{KPR}}(\eta) = \mathbf{Q} \mathbf{V} \mathbf{S}^{-1} \mathbf{W}_\eta \mathbf{U}^\top \mathbf{H} \mathbf{y}.$$

2. Proof of Propositions 3.1 and 3.2. Recall $\widehat{\beta}_j^w(h_j) = \beta_j^w - \widehat{B}_j(h_j)$ and $\mathbf{A} = \mathbf{Q} \mathbf{V} \mathbf{W} \mathbf{S}^{-1} \mathbf{U}^\top \mathbf{H} \mathbf{L}_\psi^\top$. Plugging in the definition of β_j^w and $\widehat{B}_j(h_j)$, we get

$$\begin{aligned}\widehat{\beta}_j^w(h_j) &= \sum_{m \neq j} \xi_{jm}^w \beta_m^* + \xi_{jj}^w \beta_j^* - \sum_{m \neq j} \xi_{jm}^w \beta_m^{\text{init}} - h_j (\xi_{jj}^w - 1) \beta_j^{\text{init}} + z_j^w \\ &= ((1 - h_j) \xi_{jj}^w + h_j) \beta_j^* + \sum_{m \neq j} \xi_{jm}^w (\beta_m^* - \beta_m^{\text{init}}) + \\ (3) \quad &h_j (\xi_{jj}^w - 1) (\beta_j^* - \beta_j^{\text{init}}) + z_j^w,\end{aligned}$$

where $z_j^w = \sum_{i=1}^n a_{ji} \tilde{\epsilon}_i$. To prove the asymptotic normality of z_j^w , we first check Lindeberg's condition; that is,

$$\lim_{n \rightarrow \infty} \frac{1}{s_{n,j}^2} \sum_{i=1}^n \mathbb{E} [a_{ji}^2 \tilde{\epsilon}_i^2 \times I \{|a_{ji} \tilde{\epsilon}_i| > t s_{n,j}\}] = 0, \text{ for all } t > 0,$$

where $s_{n,j}^2 = \sum_{i=1}^n a_{ji}^2$. Let ε be a random variable distributed like every $\tilde{\varepsilon}_i$. Then,

$$\begin{aligned} \frac{1}{s_{n,j}^2} \sum_{i=1}^n \mathbb{E} [a_{ji}^2 \tilde{\varepsilon}_i^2 \times I\{|a_{ji} \tilde{\varepsilon}_i| > ts_{n,j}\}] &= \frac{1}{s_{n,j}^2} \sum_{i=1}^n \{a_{ji}^2 \mathbb{E} [\varepsilon^2 \times I\{|\varepsilon| > ts_{n,j} |a_{ji}|^{-1}\}]\} \\ &\leq \frac{1}{s_{n,j}^2} \left(\sum_{i=1}^n a_{ji}^2 \right) \mathbb{E} \left[\varepsilon^2 \times I \left\{ |\varepsilon| > ts_{n,j} \left\{ \max_{i=1,\dots,n} |a_{ji}| \right\}^{-1} \right\} \right] \\ &= \mathbb{E} \left[\varepsilon^2 \times I \left\{ |\varepsilon| > \frac{t \sqrt{\sum_{i=1}^n a_{ji}^2}}{\max_{i=1,\dots,n} |a_{ji}|} \right\} \right]. \end{aligned}$$

Since

$$\lim_{n \rightarrow \infty} \frac{\max_{i=1,\dots,n} |a_{ji}|}{\sqrt{\sum_{i=1}^n a_{ji}^2}} = 0,$$

by using the dominated convergence theorem, we get

$$\lim_{n \rightarrow \infty} \frac{1}{s_{n,j}^2} \sum_{i=1}^n \mathbb{E} [a_{ji}^2 \tilde{\varepsilon}_i^2 \times I\{|a_{ji} \tilde{\varepsilon}_i| > ts_{n,j}\}] = 0 \text{ for all } t > 0.$$

Next, using the Lindeberg central limit theorem, we get $s_{n,j}^{-1} z_j^w \xrightarrow{d} N(0, 1)$ as $n \rightarrow \infty$. Lastly, we find the explicit form of $s_{n,j}$ by noting that

$$s_{n,j}^2 = \sum_{i=1}^n a_{ji}^2 = (\mathbf{A}\mathbf{A}^\top)_{(j,j)} = \Omega_{jj}^w,$$

which completes the proof of Proposition 3.1.

Next, we prove Proposition 3.2. Recall that $R_{jj}^w = \sigma^2 \{\mathbf{Q}\mathbf{V}\mathbf{W}^2\mathbf{S}^{-2}\mathbf{V}^\top\mathbf{Q}\}_{(j,j)}$. Then, we can write

$$\Omega_{jj}^w = R_{jj}^w + \mathbf{e}_j^\top \mathbf{Q}\mathbf{V}\mathbf{W}\mathbf{S}^{-1}\mathbf{U}^\top \mathbf{H}^{1/2} (\mathbf{H}^{1/2} \mathbf{L}_\psi^\top \mathbf{L}_\psi \mathbf{H}^{1/2} - \sigma^2 \mathbf{I}_n) \mathbf{H}^{1/2} \mathbf{U}\mathbf{S}^{-1} \mathbf{W}\mathbf{V}^\top \mathbf{Q} \mathbf{e}_j,$$

where $\mathbf{e}_j \in \mathbb{R}^n$ is the vector with one in its j -th entry and zero elsewhere. By Assumption (A1), we know that $\|\mathbf{H}^{1/2} \mathbf{L}_\psi^\top \mathbf{L}_\psi \mathbf{H}^{1/2} - \sigma^2 \mathbf{I}_n\|_2 = \|\mathbf{L}_\psi \mathbf{H} \mathbf{L}_\psi^\top - \sigma^2 \mathbf{I}_n\|_2 = o(1)$ as $n \rightarrow \infty$. Thus, we have

$$\mathbf{e}_j^\top \mathbf{Q}\mathbf{V}\mathbf{W}\mathbf{S}^{-1}\mathbf{U}^\top \mathbf{H}^{1/2} (\mathbf{H}^{1/2} \mathbf{L}_\psi^\top \mathbf{L}_\psi \mathbf{H}^{1/2} - \sigma^2 \mathbf{I}_n) \mathbf{H}^{1/2} \mathbf{U}\mathbf{S}^{-1} \mathbf{W}\mathbf{V}^\top \mathbf{Q} \mathbf{e}_j = o(R_{jj}^w).$$

This further leads to

$$\frac{\Omega_{jj}^w}{R_{jj}^w} = 1 + o(1), \text{ as } n \rightarrow \infty.$$

Since $z_j^w / \sqrt{\Omega_{jj}^w} \xrightarrow{d} N(0, 1)$, we have

$$\frac{z_j^w}{\sqrt{R_{jj}^w}} = \frac{z_j^w}{\sqrt{\Omega_{jj}^w}} \times \sqrt{\frac{\Omega_{jj}^w}{R_{jj}^w}} \xrightarrow{d} N(0, 1) \text{ as } n \rightarrow \infty.$$

This completes the proof of Proposition 3.2.

3. Proof of Theorem 3.3. Recall that $\check{\mathbf{X}} = \mathbf{H}^{1/2} \mathbf{X} \mathbf{D} \mathbf{\Delta}^{1/2}$ and $\tilde{\boldsymbol{\beta}}^* = \mathbf{D}^\top \boldsymbol{\beta}^*$. The next lemma characterizes $\left\| \tilde{\boldsymbol{\beta}}(\lambda) - \tilde{\boldsymbol{\beta}}^* \right\|_1$, where $\tilde{\boldsymbol{\beta}}(\lambda)$ is given in eq. (13) in the main text.

LEMMA 3.1. *Suppose that $\|\mathbf{Q}\|_2 = 1$ and each column of \mathbf{X} has been scaled so that $\|\mathbf{X} \mathbf{d}_j\|_{\mathbf{H}}^2 = n$ for $j = 1, \dots, p$, where \mathbf{d}_j is the j -th eigenvector of \mathbf{Q} . For any $c_0 > 0$, if*

$$\lambda = 2 \sqrt{2c^* n \log p (1 + c_0) \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2},$$

then under Assumptions (A1)–(A3), we have

$$\left\| \tilde{\boldsymbol{\beta}}(\lambda) - \tilde{\boldsymbol{\beta}}^* \right\|_1 = O_p \left(s_0 \sqrt{n^{-1} \log p} \right) \text{ as } n \rightarrow \infty.$$

PROOF OF LEMMA 3.1. We first define some additional notations. For any index set $A \subset \{1, \dots, p\}$, let $\check{\mathbf{X}}_A$ denote the submatrix of $\check{\mathbf{X}}$ with columns indexed by A , and let \mathbf{P}_A denote the projection matrix onto the column space of $\check{\mathbf{X}}_A$. Define

$$\chi_m^* = \max_{|A|=m, s \in \{\pm 1\}^m} \left| \boldsymbol{\varepsilon}^\top \frac{\check{\mathbf{X}}_A (\check{\mathbf{X}}_A' \check{\mathbf{X}}_A)^{-1} \mathbf{s} \lambda - (\mathbf{I} - \mathbf{P}_A) \check{\mathbf{X}} \mathbf{\Delta}^{-1/2} \tilde{\boldsymbol{\beta}}^*}{\left\| \check{\mathbf{X}}_A (\check{\mathbf{X}}_A' \check{\mathbf{X}}_A)^{-1} \mathbf{s} \lambda - (\mathbf{I} - \mathbf{P}_A) \check{\mathbf{X}} \mathbf{\Delta}^{-1/2} \tilde{\boldsymbol{\beta}}^* \right\|_2} \right|,$$

where $\boldsymbol{\varepsilon} = \mathbf{H}^{1/2} \boldsymbol{\epsilon}$, and $\lambda > 0$ is the tuning parameter. For all $m_0 \geq 0$, define the following Borel set

$$\Omega_{m_0} \equiv \{ \chi_m^* \leq t_m \ \forall m \geq m_0 \},$$

where

$$t_m = \sqrt{2 \log p (m \vee 1) (1 + c_0) \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2}$$

for some constant $c_0 > 0$. Let

$$r_1 \equiv \left(\frac{c^* \eta_1 n}{s_0 \lambda} \right)^{1/2}, \quad r_2 \equiv \left(\frac{c^* \eta_1^2 n}{s_0 \lambda^2} \right)^{1/2} \quad \text{and} \quad C \equiv \frac{c^*}{c_*}.$$

It can be checked that with the η_1 specified in Assumption (A2) and the λ specified in Lemma 3.1, r_1 and r_2 are both $O(1)$ quantities. Let $M_1^* = 2 + 4r_1^2 + 4\sqrt{C}r_2 + 4C$, $\tilde{S}_0 = \{j : \tilde{\beta}_j(\lambda) \neq 0\}$ and $S_1 = \tilde{S}_0 \cup S_0$. We follow the proof in Zhang et al. (2008), and summarize the following three major steps.

1. According to the proof of Theorem 1 in Zhang et al. (2008) (see Equation (5.26) in Zhang et al. (2008) and its derivations), we have $|S_1| \leq M_1^* s_0$ while conditioning on the event Ω_{s_0} .

2. In the second step, we show that the event Ω_{s_0} happens with high probability. In Zhang et al. (2008), this is proved for the setting where $\boldsymbol{\varepsilon} \sim N(0, \sigma^2 \mathbf{I}_n)$. However, in our case, $\boldsymbol{\varepsilon}$ is not necessarily Gaussian, and its covariance matrix, $\mathbf{H}^{1/2} \mathbf{L}_\psi^\top \mathbf{L}_\psi \mathbf{H}^{1/2}$, is not necessarily equal to \mathbf{I}_n (up to a constant). To prove this, we first notice that for all s_0 ,

$$(4) \quad \mathbb{P}(\Omega_{s_0}^c) \leq \mathbb{P}(\Omega_0^c) \leq \sum_{m=0}^{\infty} 2^{m \vee 1} \binom{p}{m} p_m,$$

where

$$p_m = \mathbb{P} \left(\left| \boldsymbol{\varepsilon}^\top \frac{\check{\mathbf{X}}_A (\check{\mathbf{X}}_A' \check{\mathbf{X}}_A)^{-1} \mathbf{s} \lambda - (\mathbf{I} - \mathbf{P}_A) \check{\mathbf{X}} \mathbf{\Delta}^{-1/2} \tilde{\boldsymbol{\beta}}^*}{\left\| \check{\mathbf{X}}_A (\check{\mathbf{X}}_A' \check{\mathbf{X}}_A)^{-1} \mathbf{s} \lambda - (\mathbf{I} - \mathbf{P}_A) \check{\mathbf{X}} \mathbf{\Delta}^{-1/2} \tilde{\boldsymbol{\beta}}^* \right\|_2} \right| \geq t_m \right)$$

with $|A| = m$ and some $\mathbf{s} \in \{\pm 1\}^m$. Next, we find an upper bound for p_m . Since $\varepsilon = \mathbf{H}^{1/2} \mathbf{L}_\psi^\top \tilde{\varepsilon}$, where $\tilde{\varepsilon}$ has *i.i.d.* sub-Gaussian entries with mean 0 and variance 1, using the Hoeffding's inequality (Wainwright, 2019), we get

$$(5) \quad p_m \leq 2 \exp\left\{-\frac{t_m^2}{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2}\right\} = 2 \exp(-(m \vee 1)(1 + c_0) \log p).$$

Combining (4) and (5), we have

$$(6) \quad \begin{aligned} \mathbb{P}(\Omega_{s_0}^c) &\leq \sum_{m=0}^{\infty} 2^{(m \vee 1)+1} \binom{p}{m} (p^{-(1+c_0)})^{m \vee 1} \\ &\leq \frac{4}{p^{1+c_0}} + 2 \sum_{m=1}^{\infty} 2^m \binom{p}{m} p^{-m(1+c_0)} \\ &= \frac{4}{p^{1+c_0}} + 2 \left((1 + 2p^{-(1+c_0)})^p - 1 \right) \\ &\leq \frac{4}{p^{1+c_0}} + 2 \left(\exp(2p^{-c_0}) - 1 \right); \end{aligned}$$

here, we use the binomial theorem and the fact that $\lim_{x \rightarrow 0} (1+x)^{1/x} = e$. Therefore, we have $\mathbb{P}(\Omega_{s_0}) = 1 - \mathbb{P}(\Omega_{s_0}^c) \rightarrow 1$ as $p \rightarrow \infty$.

3. We use the techniques in the proof of Theorem 3 in Zhang et al. (2008) to bound $\|\tilde{\beta}(\lambda) - \tilde{\beta}^*\|_1$ with the pre-specified η_1 and λ . Recall that on the event Ω_{s_0} , we have $|S_1| \leq M_1^* s_0$. Let $\check{\mathbf{X}}_1$ and $\check{\mathbf{X}}_2$, respectively, denote the the submatrix of $\check{\mathbf{X}}$ with columns indexed by S_1 and S_1^c . Similarly, let $\tilde{\beta}_1^*$ and $\tilde{\beta}_2^*$, respectively, denote the subvector of $\tilde{\beta}^*$ with entries indexed by S_1 and S_1^c ; $\tilde{\beta}_1(\lambda)$ and $\tilde{\beta}_2(\lambda)$, respectively, denote the subvector of $\tilde{\beta}(\lambda)$ with entries indexed by S_1 and S_1^c ; $\check{\mathbf{\Delta}}_1$ and $\check{\mathbf{\Delta}}_2$, respectively, denote the submatrix of $\check{\mathbf{\Delta}}$ with both rows and columns indexed by S_1 and S_1^c . Since $q^* \geq M_1^* s_0 + 1$, by Assumption (A3), the vector $\mathbf{v}_1 = \check{\mathbf{X}}_1 \check{\mathbf{\Delta}}_1^{-1/2} \left(\tilde{\beta}_1(\lambda) - \tilde{\beta}_1^* \right)$ satisfies

$$(7) \quad \|\mathbf{v}_1\|_2^2 \geq c_* n \left\| \check{\mathbf{\Delta}}_1^{-1/2} \left(\tilde{\beta}_1(\lambda) - \tilde{\beta}_1^* \right) \right\|_2^2.$$

Denoting by \mathbf{P}_1 the projection matrix onto the column space of $\check{\mathbf{X}}_1$, we have

$$(8) \quad \begin{aligned} \|\mathbf{v}_1\|_2 &\leq \|\mathbf{v}_1 + \mathbf{P}_1(\tilde{\mathbf{y}} - \check{\mathbf{X}}_1 \check{\mathbf{\Delta}}_1^{-1/2} \tilde{\beta}_1(\lambda))\|_2 + \|\mathbf{P}_1(\tilde{\mathbf{y}} - \check{\mathbf{X}}_1 \check{\mathbf{\Delta}}_1^{-1/2} \tilde{\beta}_1(\lambda))\|_2 \\ &\leq \|\mathbf{P}_1 \check{\mathbf{X}}_2 \check{\mathbf{\Delta}}_2^{-1/2} \tilde{\beta}_2^* + \mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi \varepsilon\|_2 + \|\mathbf{P}_1(\tilde{\mathbf{y}} - \check{\mathbf{X}}_1 \check{\mathbf{\Delta}}_1^{-1/2} \tilde{\beta}_1(\lambda))\|_2 \\ &\leq \|\check{\mathbf{X}}_2 \check{\mathbf{\Delta}}_2^{-1/2} \tilde{\beta}_2^*\|_2 + \|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi \varepsilon\|_2 + n^{-1/2} \|\Sigma_{11}^{-1/2} \mathbf{g}_1\|_2, \end{aligned}$$

where $\Sigma_{11} \equiv n^{-1} \check{\mathbf{X}}_1^\top \check{\mathbf{X}}_1$, $\tilde{\mathbf{y}} = \mathbf{H}^{1/2} \mathbf{y}$ and $\mathbf{g}_1 = \check{\mathbf{X}}_1^\top \left(\tilde{\mathbf{y}} - \check{\mathbf{X}}_1 \check{\mathbf{\Delta}}_1^{-1/2} \tilde{\beta}_1(\lambda) \right)$. Note that the KKT conditions of eq. (13) in the main text yield that $\|\mathbf{g}_1\|_\infty \leq \lambda$. Then, by Assumption (A3), we have

$$\|\Sigma_{11}^{-1/2} \mathbf{g}_1\|_2 \leq c_*^{-1/2} \|\mathbf{g}_1\|_\infty \sqrt{|S_1|} \leq \lambda (M_1^* s_0 / c_*)^{1/2}.$$

Also, since $S_1^c \subset S_0^c$, $\|\mathbf{X} \mathbf{d}_j\|_{\mathbf{H}}^2 = n$, and $\|\mathbf{Q}\|_2 = 1$, we have $\|\check{\mathbf{X}}_2 \check{\mathbf{\Delta}}_2^{-1/2} \tilde{\beta}_2^*\|_2 \leq \sqrt{n} \eta_1$. We next find an upper bound for $\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi \varepsilon\|_2$. Using the Hanson-Wright inequality (Theorem 2.1 in Rudelson et al., 2013), we have

$$(9) \quad \mathbb{P} \left(\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top \tilde{\varepsilon}\|_2 \geq \|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top\|_F + t \right) \leq \exp \left\{ -\frac{ct^2}{\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top\|_2^2} \right\},$$

for some constant c . By properties of matrix norms, we have

$$\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top\|_2^2 \leq \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2; \|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top\|_F^2 \leq \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 |S_1|;$$

here, we use the facts that $\|\mathbf{P}_1\|_2 = 1$ and $\|\mathbf{P}_1\|_F = |S_1|$. Thus,

$$\mathbb{P} \left(\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top \tilde{\boldsymbol{\epsilon}}\|_2 \geq \sqrt{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 |S_1|} + t \right) \leq \exp \left\{ -\frac{ct^2}{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2} \right\}.$$

Letting $t^2 = \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 |S_1| \log p$, there exists a Borel set $\tilde{\Omega}$ with $\mathbb{P}(\tilde{\Omega}) \rightarrow 1$ as $p \rightarrow \infty$, such that on this set

$$(10) \quad \|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top \tilde{\boldsymbol{\epsilon}}\|_2 \leq \sqrt{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 |S_1|} \left(1 + \sqrt{\log p} \right).$$

Therefore, on the event $\Omega_{s_0} \cap \tilde{\Omega}$, we have

$$\|\mathbf{P}_1 \mathbf{H}^{1/2} \mathbf{L}_\psi^\top \tilde{\boldsymbol{\epsilon}}\|_2 \leq \sqrt{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 M_1^* s_0} \left(1 + \sqrt{\log p} \right).$$

Thus,

$$(11) \quad \begin{aligned} \left\| \boldsymbol{\Delta}_1^{-1/2} \left(\tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^* \right) \right\|_2 &\leq (c_* n)^{-1/2} \|\mathbf{v}_1\|_2 \\ &\leq c_*^{-1/2} \eta_1 + n^{-1/2} \sqrt{\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 M_1^* s_0} \left(1 + \sqrt{\log p} \right) \\ &\quad + n^{-1/2} \lambda \left(\frac{M_1^* s_0}{n c_*} \right)^{1/2}. \end{aligned}$$

Since $\lambda = 2\sqrt{2c_* n \log p (1 + c_0)} \|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2$, $\|\mathbf{L}_\psi^\top \mathbf{H} \mathbf{L}_\psi\|_2 = \sigma^2 + o(1)$, $\eta_1 = O(\sqrt{n^{-1} s_0 \log p})$ and $M_1^* = O(1)$, we have

$$\left\| \tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^* \right\|_2 \leq \left\| \boldsymbol{\Delta}_1^{1/2} \right\|_2 \left\| \boldsymbol{\Delta}_1^{-1/2} \left(\tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^* \right) \right\|_2 = O_p \left(\sqrt{n^{-1} s_0 \log p} \right).$$

Therefore,

$$\|\tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^*\|_1 \leq \sqrt{s_0} \|\tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^*\|_2 = O_p \left(s_0 \sqrt{n^{-1} \log p} \right).$$

Finally,

$$\|\tilde{\boldsymbol{\beta}}(\lambda) - \tilde{\boldsymbol{\beta}}^*\|_1 \leq \|\tilde{\boldsymbol{\beta}}_1(\lambda) - \tilde{\boldsymbol{\beta}}_1^*\|_1 + \eta_1 = O_p \left(s_0 \sqrt{n^{-1} \log p} \right).$$

This completes the proof. \square

Now we use Lemma 3.1 to prove Theorem 3.3. First, note that since $s_0 = o((n/\log p)^r)$ for some $r \in (0, 1/2)$, we have

$$(12) \quad \left\| \tilde{\boldsymbol{\beta}}^* - \tilde{\boldsymbol{\beta}}(\lambda) \right\|_1 = o_p \left\{ \left(\frac{\log p}{n} \right)^{1/2-r} \right\}.$$

Thus,

$$(13) \quad \begin{aligned} |\zeta_j(h_j)| &= \left| \sum_{m=1}^p \xi_{jm}^w (\beta_m^* - \beta_m^{init}) - ((1 - h_j) \xi_{jj}^w + h_j) (\beta_j^* - \beta_j^{init}) \right| \\ &= \left| [(\mathbf{QVWV}^\top - (1 - h_j)\boldsymbol{\Xi} - h_j \mathbf{I}_p) (\boldsymbol{\beta}^* - \boldsymbol{\beta}^{init})]_j \right| \\ &\leq \left\| [(\mathbf{QVWV}^\top - (1 - h_j)\boldsymbol{\Xi} - h_j \mathbf{I}_p) \mathbf{D}]_{(j,\cdot)} \right\|_\infty \left\| \tilde{\boldsymbol{\beta}}^* - \tilde{\boldsymbol{\beta}}(\lambda) \right\|_1. \end{aligned}$$

Combining (12) and (13), it can be seen that

$$\lim_{n \rightarrow \infty} \Pr \left\{ |\zeta_j(h_j)| \leq \left\| [(\mathbf{Q}\mathbf{V}\mathbf{W}\mathbf{V}^\top - (1-h_j)\mathbf{\Xi} - h_j\mathbf{I}_p)\mathbf{D}]_{(j,\cdot)} \right\|_\infty \left(\frac{\log p}{n} \right)^{1/2-r} \right\} = 1.$$

Finally, Theorem 3.3 directly follows from Propositions 3.1 and 3.2.

4. Proof of Theorem 3.4. We first note that $P_j^w(h_j) \leq \alpha$ is equivalent to

$$\left| \widehat{\beta}_j^w(h_j) \right| \geq \Psi_j(h_j) + q_{(1-\alpha/2)} \sqrt{R_{jj}^w}.$$

Since $\left| \widehat{\beta}_j^w(h_j) \right| = \left| (1-h_j)\xi_{jj}^w + h_j \beta_j^* + \zeta_j(h_j) + z_j^w \right|$, we have

$$\begin{aligned} & \Pr \left\{ \left| \widehat{\beta}_j^w(h_j) \right| \geq \Psi_j(h_j) + q_{(1-\alpha/2)} \sqrt{R_{jj}^w} \right\} \geq \\ & \Pr \left\{ \left| (1-h_j)\xi_{jj}^w + h_j \beta_j^* \right| - |\zeta_j(h_j)| - |z_j^w| \geq \Psi_j(h_j) + q_{(1-\alpha/2)} \sqrt{R_{jj}^w} \right\}. \end{aligned}$$

Hence, it suffices to show that as $n \rightarrow \infty$,

$$(14) \quad \Pr \left\{ \left| (1-h_j)\xi_{jj}^w + h_j \beta_j^* \right| - |\zeta_j(h_j)| - |z_j^w| \geq \Psi_j(h_j) + q_{(1-\alpha/2)} \sqrt{R_{jj}^w} \right\} \geq \psi.$$

Since $(R_{jj}^w)^{-1/2} z_j^w \xrightarrow{d} N(0, 1)$ as $n \rightarrow \infty$, if

$$(15) \quad \frac{\left| \left((1-h_j)\xi_{jj}^w + h_j \beta_j^* \right) - |\zeta_j(h_j)| - \Psi_j(h_j) - q_{(1-\alpha/2)} \sqrt{R_{jj}^w} \right|}{\sqrt{R_{jj}^w}} \geq q_{(1-\psi/2)},$$

then (14) holds. Since $\lim_{n \rightarrow \infty} \Pr(|\zeta_j(h_j)| \leq \Psi_j(h_j)) = 1$ (Theorem 3.3), we know that as $n \rightarrow \infty$, (15) holds if

$$|\beta_j^*| \geq |(1-h_j)\xi_{jj}^w + h_j|^{-1} \left(2\Psi_j(h_j) + (q_{(1-\alpha/2)} + q_{(1-\psi/2)}) \sqrt{R_{jj}^w} \right);$$

this completes the proof of Theorem 3.4.

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